

## Research



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## Causal learning, counterfactual reasoning and pretend play: a cross-cultural comparison of Peruvian, mixed- and low-socioeconomic status U.S. children

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Pretend play universally emerges during early childhood and may support the development of causal inference and counterfactual reasoning. However, the amount of time spent pretending, the value that adults place on pretense and the scaffolding adults provide varies by both culture and socioeconomic status (SES). In middle class U.S. preschoolers, accuracy on a pretense-based causal reasoning task predicted performance on a similar causal counterfactual task. We explored the relationship between cultural environment, pretense and counterfactual reasoning in low-income Peruvian ( $N = 62$ ) and low-income U.S. ( $N = 57$ ) 3- to 4-year olds, and contrast findings against previous findings in an age matched, mixed-SES U.S. sample ( $N = 60$ ). Children learned a novel causal relationship, then answered comparable counterfactual and pretense-based questions about the relationship. Children's responses for counterfactual and pretense measures differed across populations, with Peruvian and lower income U.S. children providing fewer causally consistent responses when compared to middle class U.S. children. Nevertheless, correlations between the two measures emerged in all populations. Across cohorts, children also provided more causally consistent answers during pretense than counterfactually. Findings strengthen the hypothesis that causal pretend play is related to causal counterfactual reasoning across cultural contexts, while also suggesting a role for systematic environmental differences.

This article is part of the theme issue 'Thinking about possibilities: mechanisms, ontogeny, functions and phylogeny'.

## 1. Causal learning, counterfactual reasoning, and pretend play: a cross-cultural comparison of Peruvian, mixed- and low-socioeconomic status U.S. children

Children from a variety of backgrounds spontaneously engage in pretend play<sup>1</sup>. Pretend play universally emerges in typically developing human children between 18 and 24 months of age, and the frequency and complexity of pretense increases during early childhood (for reviews see [3,4]). Even children from cultures where adults discourage pretense spontaneously pretend [3,5], suggesting that pretense may have an innate and evolutionary basis. This may seem perplexing, given that there are no obvious adaptive benefits of pretending. In fact, superficially, it would seem a waste of time to engage in a fictional world when children still have so much to learn about the real one.

However, some researchers have argued that pretend play scaffolds other cognitive skills (e.g. [6]) including theory of mind, divergent thinking and counterfactual reasoning<sup>2</sup>. If these hypotheses hold, pretense would not only be

64 universal and innate, but arguably provide adaptive benefits.  
 65 Specifically, one role for pretense might be in facilitating  
 66 counterfactual causal reasoning. By contrast, other research-  
 67 ers have argued that extensive pretend play may not play a  
 68 significant, or culturally universal, role in learning [9]. In  
 69 this paper, we leverage cross-cultural and cross-socioeco-  
 70 nomic status (SES) data to test the universality of a  
 71 relationship between causal counterfactual reasoning and  
 72 pretense in preschool-aged children.

73 While children spontaneously engage in pretense by age  
 74 2, there are mixed findings as to when children can reason  
 75 counterfactually, with some procedures showing competence  
 76 by age 4, and others not until 6 or 8 years of age or even much  
 77 older (for a recent review see [10]). While some of this  
 78 variance can likely be explained by researchers leveraging  
 79 different definitions of counterfactual reasoning (see footnote  
 80 2), the complexity of specific experimental procedures or the  
 81 use of different types of counterfactuals (see [11,12] for  
 82 further discussion), in all cases counterfactual competence  
 83 appears to emerge later than pretense abilities.

84 A number of experiments have reliably shown that 4-year  
 85 olds can accurately reason counterfactually about determinis-  
 86 tic physical causal systems ([11,13,14]; see [15] for related  
 87 results with older children). For example, both Buchsbaum  
 88 *et al.* [13] and Nyhout *et al.* [11] used 'blicket machine' style  
 89 tasks (see [16]) where children were introduced to a novel  
 90 machine and learned to activate it by placing a specific  
 91 block on the machine. Then children answered counterfactual  
 92 questions about this newly learned physical causal relation-  
 93 ship. Nyhout *et al.* [11] found that, in a 'blicket machine'  
 94 task, preschoolers were able to correctly reason about overde-  
 95 termined scenarios, where an outcome has multiple causes  
 96 (e.g. would the machine still have activated if only one of  
 97 two 'blickets' was removed), despite struggling with overde-  
 98 termined counterfactual scenarios until much later ages in  
 99 other contexts (e.g. [11,12,17]).

100 Most relevant to the present paper, Buchsbaum *et al.* [13]  
 101 argue that pretense and counterfactual reasoning exercise the  
 102 same underlying causal reasoning mechanisms (see [18,19]  
 103 for similar arguments) and provide correlational evidence  
 104 to support this hypothesis. As noted above, the content of  
 105 counterfactual questions appears to play an important role  
 106 in children's counterfactual competence. The content of  
 107 pretend play can also vary widely, reflecting children's differ-  
 108 ing domain knowledge. In order to help control for potential  
 109 differences in content and domain knowledge, Buchsbaum  
 110 *et al.* compared children's responses to a specific counterfac-  
 111 tual and pretense prompt in exactly the same situation—a  
 112 blicket machine-like 'birthday machine'—and so could ask  
 113 the children to produce pretense responses that were highly  
 114 comparable to the counterfactual ones.

115 In two experiments, children learned a novel causal  
 116 relationship: that a specific block, the zando, activated a  
 117 zando machine and caused it to play 'Happy Birthday'; how-  
 118 ever the other block, the non-zando, did not activate the  
 119 machine. Then children were asked counterfactual questions  
 120 about the causal relationship between the blocks and the  
 121 machine (e.g. 'What would happen if the zando was not a  
 122 zando? Would the machine play music or would it not play  
 123 music?'). Even 3- and 4-year olds can sometimes correctly  
 124 answer questions about such machines. After this, the exper-  
 125 imenter engaged children in a pretense sequence. During the  
 126 pretense phase, children were asked comparable questions

about a pretend zando machine, a pretend zando and a  
 pretend non-zando.

Buchsbaum and colleagues found that children were more  
 successful at answering hypothetical questions framed within  
 causal pretense, than matching questions framed counterfac-  
 tually, and argued that children may be more successful at  
 reasoning about alternatives to reality in a pretense context.  
 They also found a correlation between the inferences children  
 made while engaged in pretense and those that they made  
 while reasoning counterfactually. Children who, for example,  
 stated that the pretend zando would activate the pretend  
 machine, also provided correct answers to the counterfactual  
 reasoning questions. This was true even after controlling for  
 age, executive functioning, which is thought to play a role in  
 counterfactual reasoning (e.g. [8,20]), and conservation of  
 number, a cognitive task on which children improve during  
 this same time period, but which is not hypothesized to be  
 related to counterfactual reasoning, to help control for devel-  
 opmental differences in cognitive abilities not specific to  
 counterfactual reasoning or pretense. Together, this supports  
 the hypothesis that causal reasoning during pretense may  
 draw upon the same emerging cognitive capacities as counter-  
 factual causal reasoning, and perhaps that early emerging  
 pretend abilities may even support the development of later  
 counterfactual reasoning abilities.

Intuitively, such a link might exist because both pretend play  
 and counterfactual reasoning are quite similar—both require the  
 ability to 'quarantine' reality in order to reason about alternative  
 fictional possibilities and their outcomes [2,3,21–23]. For example,  
 while pretending, a child may imagine that a banana is a tele-  
 phone, then work through various instances of what would  
 happen if it were indeed a telephone rather than a banana.  
 This is structurally similar to thinking through counterfactual  
 scenarios, where a hypothetical change to a premise may bring  
 about a range of different outcomes, and an agent must reason  
 about the various possibilities.

As discussed above, researchers have found that children  
 spontaneously pretend at a much younger age, around 18  
 months, than they first demonstrate competence in counterfac-  
 tual reasoning, around age 4. In fact, the experimental  
 literature suggests that placing hypothetical questions within  
 a pretense or fantasy setting actually enhances young chil-  
 dren's ability to reason from a false premise [24–27]. In these  
 studies, pretense may have supported children's ability to  
 quarantine a fictitious premise from real-world knowledge  
 and reason through causal outcomes. Taken together this  
 could suggest that the ability to reason causally during pre-  
 tense emerges early and supports further development of  
 complex real-world counterfactual reasoning.

While pretense is ubiquitous in young children regardless of  
 culture and economic standing, empirical work suggests that  
 the sociocultural context substantially shapes children's day to  
 day pretense activities. Several cross-cultural observational  
 studies have documented variance in pretend play along several  
 dimensions. These dimensions include the amount of time that  
 children spend engaged in pretense, the value that parents and  
 caretakers place on pretend play, the extent that parents engage  
 in pretense along with their children, the types of partners  
 young children pretend with (e.g. same-aged children, adults,  
 older children and alone), the subject matter of pretense and  
 the types of props that are used while pretending [5,28–39].

In one study, Callaghan *et al.* [28] sampled Indian, Peru-  
 vian and Canadian mother-child dyads, and measured

127 maternal beliefs about pretense, observed mother–child  
 128 interactions, and children’s tendencies to pretend, both spon-  
 129 taneously and when interacting with an adult who was  
 130 pretending. In this study, all Canadian mothers reported pre-  
 131 tending with their children; however, only 42% of Peruvian  
 132 mothers and 24% of Indian mothers reported doing so. Call-  
 133 aghan *et al.* next measured both the spontaneous pretense  
 134 actions of children across cultures and children’s response  
 135 to an experimenter’s pretense action. North American chil-  
 136 dren performed more spontaneous pretense acts than either  
 137 the Indian and Peruvian children, and were more likely  
 138 follow the experimenter’s pretense action with a subsequent  
 139 pretense action. Furthermore, the average age when children  
 140 did so was about 1-year younger in the North American  
 141 sample (34.4 months), than in either the Indian (46.5  
 142 months) or Peruvian samples (45.8 months). In the present  
 143 paper, we leverage this work by comparing Peruvian children  
 144 to children from the U.S. on pretense and counterfactual  
 145 causal reasoning tasks.

146 The early cognitive developmental literature (though see  
 147 [40] for a critical review of early findings), as well as more  
 148 recent findings, suggest that higher SES children spend more  
 149 time pretending than low-SES children, although children  
 150 from all economic brackets pretend. For example, Doyle *et al.*  
 151 [41] found that middle class 5- to 7-year-old North American  
 152 children spent more time in social pretense and pretended  
 153 for longer periods of time when compared with low-SES  
 154 North American children. There are similar findings for  
 155 children in Brazil [36], and in Israel and South Africa [42].

156 This naturally occurring variance raises questions about  
 157 how socioeconomic and cultural differences in pretend play  
 158 may relate to other aspects of cognitive development, such  
 159 as counterfactual reasoning specifically. In the present  
 160 paper, we explore this by extending the Buchsbaum *et al.*  
 161 [13] paradigm to lower income 3- and 4-year-old children  
 162 in Peru as well as lower income 3- and 4-year-old U.S. chil-  
 163 dren enrolled in Head Start programs. These samples are  
 164 directly compared to the middle class U.S. sample from  
 165 Buchsbaum *et al.* [13], Experiment 2.

166 In addition to the findings discussed above, our own pre-  
 167 existing observations of the specific preschools we partnered  
 168 with in Peru during this research suggest that teachers and  
 169 adults may provide less scaffolding for pretense than do  
 170 adults in the U.S. For example, the traditional preschool  
 171 environment in Peru does not provide children with props  
 172 or space to pretend in, and this is commonly provided to chil-  
 173 dren in the USA, including in the lower income Head Start  
 174 preschools we also partnered with. Researchers also pre-  
 175 viously noted that in the Peruvian classrooms, more time is  
 176 spent in teacher led group activities, and less in self-directed  
 177 or choice-based free play. This could indicate that children  
 178 have fewer opportunities to enter into pretense during the  
 179 school day, and that (consistent with the [28] findings) pre-  
 180 tense is less supported by adults. Previous researchers have  
 181 found that classroom environment impacts children’s ten-  
 182 dency to engage in pretense [43].

183 This raises questions about the extent to which the Buchs-  
 184 baum *et al.* findings will replicate in the Peruvian and low-  
 185 SES U.S. samples. First, will the correlation between pretense  
 186 and counterfactual reasoning, and comparatively superior  
 187 performance on pretense versus counterfactual reasoning  
 188 tasks, replicate in a context where pretense activities may  
 189 differ? If causal reasoning during pretense and in

counterfactual scenarios is underpinned by the same  
 cognitive capacities, then the correlation should replicate  
 regardless of cultural or socioeconomic background. More-  
 over, if reasoning about imagined causal scenarios is easier  
 in pretense, perhaps facilitating later counterfactual reason-  
 ing, then we might see improved performance on questions  
 couched in a pretense scenario versus equivalent counterfac-  
 tual questions. On the other hand, if causal counterfactual  
 reasoning and similar reasoning during pretense are not  
 underpinned by the same abilities, but instead reflect some  
 other common factors that are particular to the middle class  
 U.S. children in the Buchsbaum *et al.*, sample, then we  
 could see a decoupling of this correlation in the Peruvian  
 and/or low-SES samples.

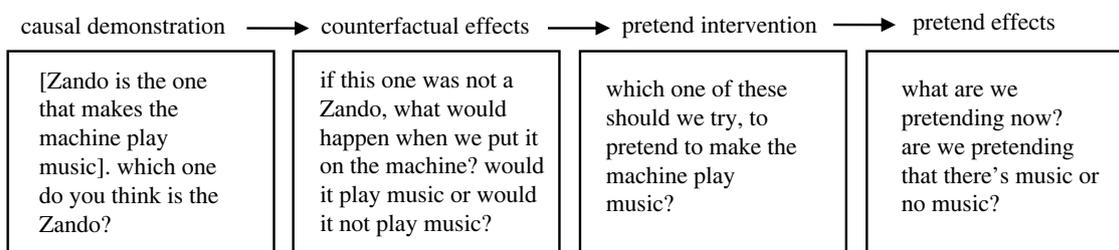
Additionally, there could be differences in developmental  
 timelines across the three cohorts in both or either skill set.  
 We do not know when low-SES U.S. children and Peruvian  
 children can reliably reason counterfactually, or reason causally  
 while pretending. Given the previous research, it is possible  
 that the Peruvian and low-SES children engage in less pretense  
 than the middle class North American children, and that these  
 differences could be reflected in difference in performance on  
 experimental measures. In particular, we might predict that  
 the differences in everyday pretense activity would lead to  
 differences in reasoning about pretense in a more controlled  
 experimental setting. If this type of pretense reasoning is also  
 related to counterfactual reasoning, we might see differences  
 in both these types of reasoning and might expect that the  
 two types of abilities would be correlated. If these two skill  
 sets are not underpinned by the same cognitive capacities,  
 then we may only see cultural variation in the timeline for  
 the pretense-based task, as that is the task that would presu-  
 mably be more related to day to day pretense activities, but not  
 necessarily during the counterfactual task. Alternatively, cul-  
 tural and SES differences in pretend practices may have no  
 effect on the development of causal reasoning either in pretense  
 or when reasoning counterfactually. If so, we might not see any  
 differences across our samples.

## 2. Methods

The overall structure, stimuli and methods of the experiment  
 replicate those used in Buchsbaum *et al.* [13]. As in Buchsbaum  
*et al.*, we used the stroop-like day/night task [44], to control  
 for individual differences in inhibitory control. As discussed  
 above, the inhibitory control task was included because previous  
 research has found a direct relationship between executive func-  
 tioning and counterfactual reasoning. As in Buchsbaum *et al.*, we  
 also included the Piagetian conservation of number task, as a  
 measure of general cognitive development, similar to numerical  
 age but skill based rather than age based. This was meant to con-  
 trol for non-specific improvement across cognitive tasks and  
 general cognitive development, and for factors like the willing-  
 ness and motivation to engage in experimental tasks, which  
 should affect these tasks as well as our target tasks. These two  
 measures combined allow us to partial out potential covariance  
 unrelated to the direct relationship between causal reasoning  
 during pretense and causal counterfactual reasoning.

### (a) Participants

Participants included 62 3- and 4-year-old Peruvian children, and  
 57 low-SES U.S. children. This sample size was similar to that  
 used by Buchsbaum *et al.*’s [13] Experiment 2,  $N = 60$  3- and 4-



**Figure 1.** Overview of the Monkey's birthday tasks (full script in the electronic supplementary material, table S1).

year olds. In Peru, the mean age was 4.34 (range = 3.38 to 4.99); for the low-SES U.S. children the mean age was 4.25 (range = 3.26 to 4.97). In Buchsbaum *et al.*'s Experiment 2, children's mean age was 3.89 (range = 2.71–4.95). Thus, the Peruvian and low-SES samples averaged slightly (but significantly) older than the Buchsbaum *et al.*'s Experiment 2 sample (details given in the results section). This slight age difference could have caused the Peruvian and low-SES U.S. children to score higher on experimental measures, but this would lead to the opposite pattern than the hypotheses outlined in the introduction of this paper. An additional six children in the Buchsbaum *et al.*, sample, and four children in the low-SES sample failed to learn the causal relationship in the causal demonstration phase of the pretense task even after a second demonstration and were excluded from analysis.

### (b) Demographics of Peruvian children

Peruvian children were recruited and tested at Innova schools in Lima, Peru. Innova schools is a chain of private schools that caters to families from the emerging middle class. Schools are frequently located in lower income neighbourhoods in central Lima, or in the outskirts of the city. Most families are internal immigrants from the Andean highlands and come from indigenous backgrounds. Many families own small businesses, and few parents are college educated. Families, however, are financially upwardly mobile and heavily invested in their children's education; families at Innova schools spend around 25% of their income on their children's education. See electronic supplementary material for additional demographic information for the Peruvian cohort.

### (c) Demographics of low-SES U.S. children

Low-income U.S. children were recruited and tested in Head Start schools in the San Francisco Bay area. To enroll in Head Start, children must be from families that are considered low income according to the federal poverty guidelines, which at time of retrieval was stated as \$26 400 for a family of four (<https://www.benefits.gov/benefit/1899> retrieved 3 November 2021).

### (d) Demographics of mixed-SES U.S. children

Buchsbaum *et al.* [13] sampled children primarily from U.S. middle and upper middle class backgrounds that reside in the San Francisco Bay Area. Children were tested in university affiliated preschools, science museums or other private preschools. According to 2019 Bay Area census data, the median annual income was \$99 406 in Alameda County and \$112 449 in San Francisco County. (<https://www.census.gov/quickfacts/fact/table/alameda-county-california-sanfrancisco-county-california/EDU685219> retrieved 3 November 2021).

There was some degree of SES variation across the Peruvian and U.S. samples in addition to cultural variation, with the Peruvian children being from a low-SES background than the U.S. mixed-SES children, although it is difficult to quantify this across countries, especially when considering families 'relative'

economic standing within their own community. When contrasting the Buchsbaum *et al.* sample to the Head Start sample, SES is a key differentiating variable.

## 3. Stimuli and procedure

Participants were tested in a private area in their school. Participants completed the 'Monkey's birthday' task, the day/night stroop task and the Piagetian conservation of number task. Task order was counterbalanced as in Buchsbaum *et al.* [13]. See electronic supplementary material for additional stimuli and counterbalancing details, and for information regarding translation of the experimental prompts from English to Spanish, and procedures to ensure comparability of experimental procedures across samples.

## 4. Monkey's birthday

### (a) Introduction

Figure 1, participants were introduced to a stuffed monkey, named Monkey, or Mono in Spanish. They were told that it was Monkey's birthday. Then the experimenter placed the monkey under the table so that the participant and the experimenter could prepare a surprise for Monkey's birthday. Next the experimenter introduced children to the 'zando' machine (or 'sando' machine in Spanish), which was similar to a blicket machine (see [16]), and was constructed from a wooden box. A doorbell was placed inside, and it played the 'Happy Birthday' song when the experimenter activated it via a button hidden under their foot. The experimenter told children that the zando machine played the song 'Happy Birthday'. They suggested that they use the machine to sing 'Happy Birthday' to Monkey.

### (b) Causal demonstration

Next, the experimenter introduced children to the 'zando' and the 'non-zando' blocks. The experimenter told children that the machine only plays 'Happy Birthday' when the zando block is placed on top of the machine, and that they need to figure out which of the two blocks is the zando. To do so, the experimenter placed each block on the machine individually two times. The machine only activated when the zando block was placed on it. Following this initial demonstration, the experimenter asked children which of the two blocks was the zando. If children answered incorrectly, the experimenter repeated the initial demonstration, then again asked which block was the zando. Children had to correctly identify the zando on this second trial before continuing the experiment. After children correctly identified the zando, the experimenter prompted participants to place each of the

253 blocks on the machine individually. The machine again only  
 254 activated when the zando block was placed on top. The  
 255 experimenter and child practiced singing 'Happy Birthday'  
 256 when the machine activated.

### 257 (c) Counterfactual causal effects

258 Next, the experimenter put the blocks back down on either  
 259 side of the machine and asked the child to answer two coun-  
 260 terfactual causal effects questions about the newly learned  
 261 causal relationship between the blocks and the machine.  
 262 These were, 'If this one (while pointing to the zando) was  
 263 NOT a zando, what would happen when we put it on the  
 264 machine?' and 'If this one (while pointing to the non-  
 265 zando) WAS a zando, what would happen when we put it  
 266 on the machine?'<sup>3</sup>. If children failed to provide an answer,  
 267 or gave an ambiguous or irrelevant answer, the experimenter  
 268 asked a forced choice question, which was, 'Would the  
 269 machine play music or would it not play music?'. Children  
 270 were scored on whether they answered these questions  
 271 consistently with the counterfactual identity of the blocks,  
 272 and counterfactual causal relationships (correct), or with the  
 273 real-world identity of the blocks and real-world causal  
 274 relationships (incorrect). After children answered these ques-  
 275 tions, the experimenter and the child placed each of the  
 276 blocks on the machine one more time and sang 'Happy Birth-  
 277 day'. Then the experimenter exclaimed that they were ready  
 278 to sing for Monkey, and they could bring Monkey back.  
 279

### 281 (d) Transition to pretense

282 A confederate immediately interjected, and apologetically  
 283 said that someone else needed to use the machine, and that  
 284 they needed to bring it to them. Then, the confederate carried  
 285 the zando machine, and the two blocks, out of the room. The  
 286 experimenter acted surprised and exclaimed, 'But we didn't  
 287 sing "Happy Birthday" to Monkey yet! What should we  
 288 do? Oh, I have an idea! Look what I found! I thought we  
 289 could PRETEND that this box is my machine! Then we can  
 290 keep playing!'.  
 291

### 292 (e) Pretense tasks

293 The experimenter took out a clear red plastic box (i.e. the pre-  
 294 tend zando machine) and placed it in the centre of the table.  
 295 The experimenter also took out two wooden triangle blocks,  
 296 one green and one yellow, and placed one on each side of the  
 297 pretend zando machine. The experimenter said, 'So we can  
 298 pretend that this box is my machine, and that this block is  
 299 a zando, and this block is not a zando (while pointing to  
 300 the blocks). Then we can still sing for Monkey! I'm going to  
 301 bring Monkey back, ok?'. The experimenter then took  
 302 Monkey out from under the table and placed him next to  
 303 the machine.  
 304

### 305 (i) Pretend intervention questions

306 Children were next asked two questions about how to pre-  
 307 tend that the blocks had the right causal powers. The  
 308 experimenter prompted children to place one of the two  
 309 blocks on the machine. To do so, they said, 'So if we're pre-  
 310 tending this is my machine, and this is a zando (while  
 311 pointing to one of the blocks), and this is not a zando  
 312 (while pointing to the other block), what should we do to pre-  
 313 tend to make the machine play music?'. If children failed to  
 314  
 315

provide an answer, or provided an ambiguous or irrelevant  
 answer, they were asked a forced choice question, 'Which  
 one of these should we try, to pretend to make it play  
 music?'. Here, children were asked to generate a causal inter-  
 vention within the pretend scenario. They could indicate the  
 block that was 'consistent' (the pretend zando) or 'inconsis-  
 tent' (the pretend non-zando) with the newly learned real-  
 world causal relationship. Children were scored as correct  
 (causally consistent) if they chose the pretend zando and  
 incorrect if they chose the pretend non-zando. Children  
 were not corrected if they chose the pretend non-zando.

### (ii) Pretend effects questions

Once children indicated a block, the experimenter placed it on  
 the machine, then asked children a pretend effects question,  
 'What are we pretending now?'. If children failed to provide  
 a response, or did not give a relevant response, the exper-  
 imenter asked a forced choice question, 'Are we pretending  
 that there's music or no music?'. In order to directly compare  
 counterfactual and pretend causal reasoning, the pretend effect  
 questions were designed to parallel the earlier questions about  
 the effects of a counterfactual change, where children were  
 asked what would happen if the counterfactual zando and  
 non-zando were placed on the machine. Then, the exper-  
 imenter suggested they try placing the other block on the  
 machine. They again asked children to indicate what they  
 should pretend was happening, and if they wanted to sing.

Following this, the experimenter suggested they pretend  
 something different. Children were instructed to reverse the  
 causal properties of the blocks (e.g. if the yellow block was  
 originally the pretend zando, then the green block became  
 the pretend zando, and the yellow block became the pretend  
 non-zando). The experimenter repeated the same set of pre-  
 tend intervention and pretend effects questions stated  
 above. In total, children were asked two pretend effects ques-  
 tions about a pretend zando and two about a pretend non-  
 zando, resulting in four total pretend effects questions. Chil-  
 dren were scored on whether they provided a pretend effect  
 response that was causally consistent or inconsistent with the  
 pretense scenario outlined by the experimenter. Children  
 were not corrected if they provided a causally inconsistent  
 response, e.g. stating that placing the pretend zando on the  
 machine did not result in pretend music.

If children respond consistently with the pretend premise,  
 they should state that the pretend zando activates the  
 machine and causes music, while the pretend non-zando  
 does not activate the machine and does not cause music. If,  
 on the other hand, children respond consistently with reality,  
 they should always state that there is no music, given that the  
 effect was pretend—neither block actually causes music to  
 play in the real world. Alternatively, if children's pretense  
 does not correlate with real-world causality, or the pretense  
 scenario set-up by the experimenter, then children should  
 answer randomly. Finally, children may prefer to pretend  
 that both objects are causally effective and always state that  
 there is music, rather than to maintain the real-world causal  
 relationship in their pretense (for instance because they do  
 not transfer the newly learned causal relationship to the pre-  
 tend scenario, or because they like the song and enjoy  
 pretending that there is music).

## 316 (f) Secondary tasks

### 317 (i) Open-ended pretense task

318 Children were asked if they wanted to pretend anything else  
319 for Monkey's Birthday. Children were allowed to pretend up  
320 to five scenarios. The experimenter and child acted out  
321 children's suggestions together.

### 322 (ii) Day/night stroop task

323 The day/night task was modelled after Gerstadt *et al.* [44].  
324 Children were instructed to say 'day' when the experimenter  
325 showed them the night card picturing a moon and stars, and  
326 'night' when the experimenter showed them the day card pic-  
327 turing a sun. There was a training phase where children had  
328 to provide four correct answers. The test phase consisted of  
329 eight of each type of card.

### 332 (iii) Conservation of number

333 Modelled after the classic task developed by Piaget [45],  
334 children were shown two rows containing five U.S. pennies  
335 a piece. The procedure began with pennies equally spaced  
336 across the two rows. The experimenter asked, 'Does this  
337 row have more coins? Does this row have more coins? Or  
338 do they both have the same number of coins?' while pointing  
339 to the appropriate row(s). Then, the rows were expanded and  
340 contracted so that one row was longer than the other; how-  
341 ever, the two still contained an equal number of pennies.  
342 The experimenter gave the same prompt. This was done  
343 one more time, resulting in a total of three prompts. Children  
344 were scored according to whether they stated that the rows  
345 contained an equal number of pennies, or if they believed  
346 that one row had more pennies than the other.

## 350 5. Results and discussion

351 The dataset from Buchsbaum *et al.*'s [13] Experiment 2 was  
352 obtained, and statistics are presented for all three samples  
353 of children. As mentioned in the participants section, mean  
354 age differed across the cohorts,  $F_{2, 176} = 14.05$ ,  $p < 0.001$ .  
355 Paired contrasts confirmed that the mixed-SES U.S. children  
356 averaged slightly younger than both Peruvian,  $t_{176} = 5.06$ ,  
357  $p < 0.001$ , and low-SES U.S. children,  $t_{176} = 3.90$ ,  $p < 0.001$ ,  
358 with no difference between the latter two groups,  $t_{176} =$   
359  $1.07$ ,  $p = 0.54$ . Below, we first examine performance on the  
360 counterfactual and pretense tasks individually, and then  
361 present analyses of the relationship between performance  
362 on these two tasks across cohorts.

### 365 (a) Causal demonstration

366 Children successfully learned the novel causal relationship—  
367 that the zando activates the machine—with 87% of Peruvian  
368 children (54/62), and 91% of low-SES U.S. children (52/57)  
369 correctly identifying the zando after the experimenter's first  
370 demonstration. This is comparable to the performance of  
371 the mixed-SES U.S. sample in Buchsbaum *et al.*'s [13] Exper-  
372 iment 2, with 88% (53/60) of children providing correct  
373 responses after the first demonstration, and suggests that chil-  
374 dren in all three cohorts understood the novel causal  
375 relationship. As in Buchsbaum *et al.*, we included all children  
376 in subsequent analyses. However, excluding children who  
377 did not correctly identify the zando after the first  
378

demonstration does not change our findings (see electronic  
supplementary material).

### (b) Counterfactual effects performance

Children answered two counterfactual effects questions, result-  
ing in an overall score ranging from 0 to 2. We examined  
children's counterfactual performance using a linear model  
with children's age and cohort (Peru, low-SES U.S. and  
mixed-SES U.S.) as predictors ([13] found that counterfactual  
performance on this task improved with age in the mixed-  
SES sample). Counterfactual performance differed by cohort,  
 $F_{2, 173} = 12.57$ ,  $p < 0.001$ , and performance on the counterfac-  
tual task improved with age,  $F_{1, 173} = 8.87$ ,  $p = 0.003$ , with no  
interaction between cohort and age,  $F_{2, 173} = 0.63$ ,  $p = 0.53$ .  
Paired contrasts comparing the cohorts confirmed that both  
children in Peru and low-SES U.S. children performed more  
poorly than mixed-SES U.S. children and didn't differ from  
each other, mixed-SES versus low-SES U.S.,  $t_{173} = 4.28$ ,  
 $p < 0.001$ , mixed-SES U.S. versus Peru,  $t_{173} = 4.48$ ,  $p < 0.001$ ,  
low-SES U.S. versus Peru,  $t_{173} = 0.25$ ,  $p = 0.97$ . Thus, despite  
the overall improvement in counterfactual performance with  
age across cohorts (figure 2b), and the Peruvian and low-SES  
children's slightly higher mean age, mixed-SES U.S. children  
provided significantly more consistent responses to counterfac-  
tual questions than did the other cohorts, figure 2a. In Peru, the  
mean counterfactual effects score was 0.90 (95% CI = 0.68–  
1.12), and in the low-SES U.S. cohort, it was 0.91 (95%  
CI = 0.68–1.15). Children in both of these cohorts were not  
correct more often than chance,  $t_{61} = 0.88$ ,  $p = 0.38$  (Peru; one  
sample  $t$ -test),  $t_{56} = 0.74$ ,  $p = 0.46$  (low-SES U.S.; one sample  
 $t$ -test). By contrast, the mean counterfactual score in the  
mixed-SES cohort was 1.43 (95% CI = 1.23–1.64). Mixed-SES  
U.S. children scored significantly above chance on the  
counterfactual questions,  $t_{59} = 4.25$ ,  $p < 0.0001$ .

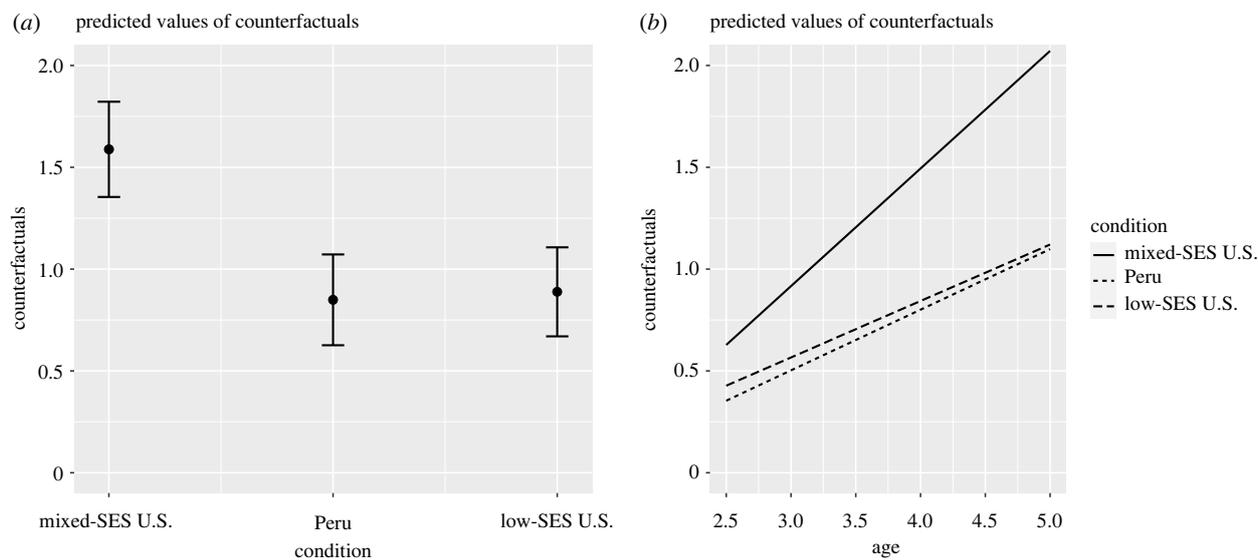
## 6. Pretense tasks

### (a) Pretend intervention

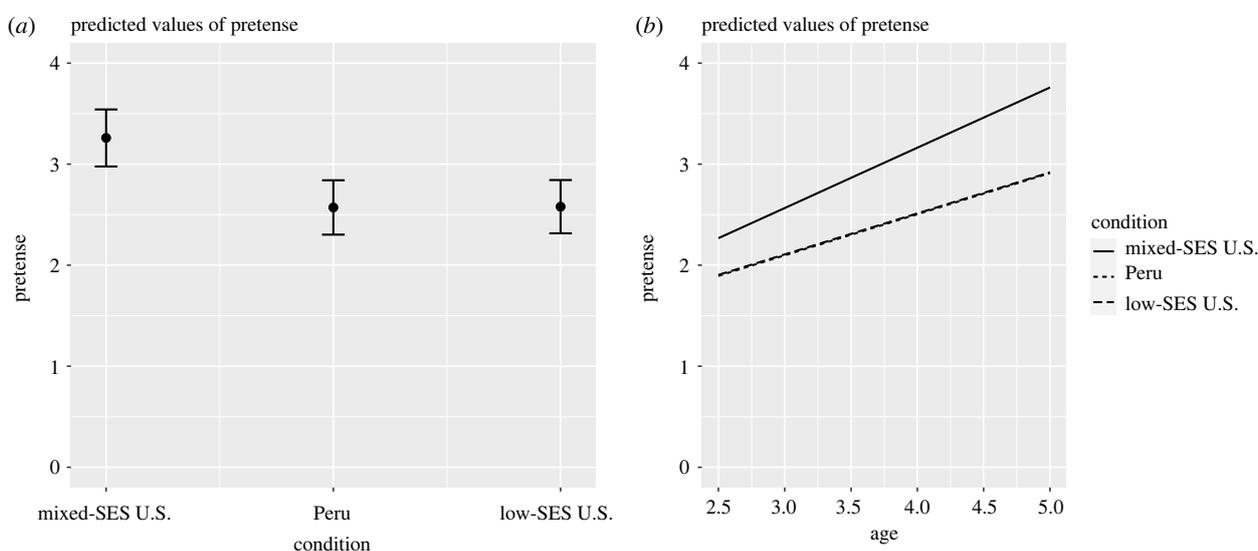
Children from all three cohorts chose to first intervene with  
the pretend zando (rather than the pretend non-zando) at  
very high rates (mixed-SES U.S.: 84%; Peru: 84%; low-SES  
U.S.: 87%), and more often than chance would predict, with  
no difference between the groups, suggesting that children  
understood the causal structure of the pretense scenario.  
More importantly, it demonstrates that children were able  
to generate a consistent causal intervention within the  
pretend context (see electronic supplementary material for  
additional details).

### (b) Pretend effects

Children answered four pretend effects questions and were  
given overall pretend effects scores of 0–4. A linear model  
with age and cohort as predictors revealed that pretend  
effects performance differed between cohorts,  $F_{1, 173} = 7.82$ ,  
 $p < 0.001$ , and performance on the pretense task improved  
with age,  $F_{1, 173} = 9.13$ ,  $p = 0.003$ , with no interaction between  
cohort and age,  $F_{1, 173} = 0.19$ ,  $p = 0.83$ . Paired contrasts com-  
paring the cohorts confirmed that both children in Peru and  
low-SES U.S. children performed more poorly than mixed-  
SES U.S. children and didn't differ from each other, mixed-  
SES versus low-SES U.S.,  $t_{173} = 3.46$ ,  $p = 0.002$ , mixed-SES



**Figure 2.** (a) Mean counterfactual score (out of 2) and confidence intervals for each cohort. (b) Counterfactual score increases with age across cohorts.



**Figure 3.** (a) Mean pretense score (out of 4) and confidence intervals for each cohort. (b) Predicted pretense score increases with age across cohorts.

versus Peru,  $t_{173} = 3.46$ ,  $p = 0.002$ , low-SES versus Peru,  $t_{173} = 0.041$ ,  $p = 0.99$ , figure 3a. Thus, as in the counterfactual task, despite the overall improvement in pretense performance with age across cohorts (figure 3b), and the Peruvian and low-SES U.S. children's slightly higher mean age, they nonetheless performed more poorly than the mixed-SES U.S. cohort collected by Buchsbaum *et al.* [13] on the causal pretense task.

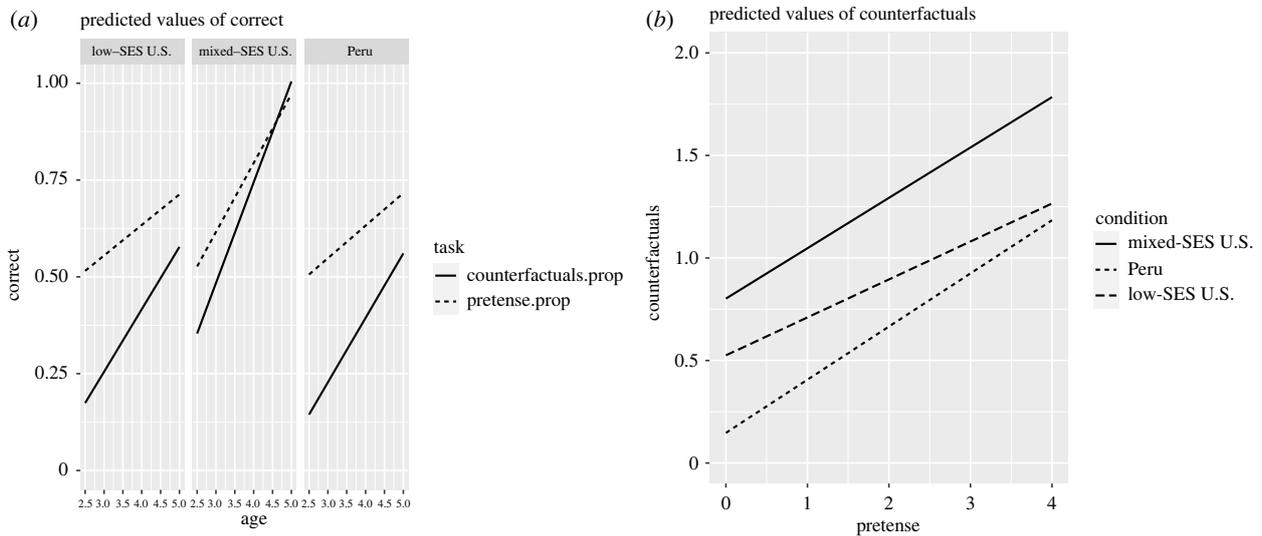
All three cohorts were significantly more likely than chance to give causally consistent pretense answers, in contrast with the counterfactual effects responses of the Peruvian and low-SES U.S. samples (Peru: mean = 2.65; 95% CI = 2.38–2.91;  $t_{61} = 4.95$ ,  $p < 0.0001$ ; low-SES: mean = 2.61; 95% CI = 2.35–2.87;  $t_{56} = 4.74$ ,  $p < 0.0001$ ; mixed-SES U.S. children: mean = 3.10; 95% CI = 2.83–3.37;  $t_{59} = 8.09$ ,  $p < 0.0001$ ). Furthermore, pretense scores did not significantly differ between the initial set of questions and when the roles of the blocks were reversed (see electronic supplementary material). Interestingly, while children in all three cohorts were above chance for both the zando and non-zando blocks, when children did provide an inconsistent answer it was more likely to be a reality-biased one—saying that the

zando block does not cause music (which is true in the real-world, but not in the pretend scenario), rather than saying that the non-zando does cause music (see electronic supplementary material).

Taken together, Peruvian and low-SES children's performance on the pretense questions suggests that they are able to reason about a counterfactual premise when it is presented in the context of pretense (whereas they were not above chance when presented with a counterfactual scenario), that they are able to maintain a newly learned a causal relationship within a pretend scenario, and that they flexibly use the pretend roles of the items in order to generate an intervention on the causal system to bring about a pretend outcome.

### (c) Comparing counterfactual effects performance and pretend effects performance

Peruvian and low-SES cohorts were successful on the pretend effects task, but at chance on the counterfactual effects task, while the mixed-SES sample was above chance on both tasks. These differences are especially notable given that



**Figure 4.** (a) Children's relative performance on the counterfactual and pretend effects tasks. (b) Higher pretense scores predict higher counterfactual scores, even with age, and performance on the day/night and conservation tasks accounted for.

performance across cohorts improved with age, and the mixed-SES U.S. sample skewed younger than the other samples. We directly compared relative performance on the pretend and counterfactual effects tasks, using a within-subjects mixed-effects linear regression on the proportion of consistent answers, with task and cohort as factors, age as a predictor, and all two-way interactions, and a random intercept for participant. There was a main effect of task,  $F_{1, 173} = 24.87$ ,  $p < 0.001$ . Children answered the pretend questions in a manner consistent with the pretend premise significantly more often than they answered the counterfactual questions consistently with the counterfactual premise. There was also a main effect of age  $F_{1, 173} = 13.46$ ,  $p < 0.001$ , and of cohort,  $F_{2, 173} = 15.98$ ,  $p < 0.001$ , with mixed-SES children outperforming the other two groups on both tasks (as discussed in previous sections), and an interaction between task and cohort  $F_{2, 173} = 3.21$ ,  $p = 0.043$ . There were no other significant interactions,  $p > 0.52$ , figure 4a.

Paired contrasts within each cohort revealed that low-SES U.S.,  $t_{175} = 3.69$ ,  $p < 0.001$ , and Peruvian children,  $t_{175} = 4.16$ ,  $p < 0.001$ , answered more consistently on the pretend effects questions than on the counterfactual effects questions. This supports the idea that, across cohorts and particularly in the low-SES and Peruvian cohorts, reasoning about an alternative to reality may be easier within pretense than when presented hypothetically. By contrast, while a qualitatively similar trend was present in the mixed-SES U.S. children, it was not significant,  $t_{175} = 0.64$ ,  $p = 0.52$ , perhaps because their performance on both tasks was already quite high (figure 4a). We will return to this point in the general discussion.

#### (d) Relationship between counterfactual and pretense measures

Finally, we explored whether children who answer the counterfactual effects questions correctly also provide more consistent responses to the pretend effects<sup>4</sup> questions. We used a linear model to explore if pretend effects score predicts children's counterfactual effects score, even when age, day/night score and conservation score are accounted for, and whether this differs across cohorts. We also included interactions between each variable and cohort, to account for

age differences between cohorts, and for any differences in performance across the secondary tasks<sup>5</sup>. Both pretend effects score  $F_{1, 155} = 13.16$ ,  $p < 0.001$ , and cohort  $F_{2, 155} = 6.10$ ,  $p = 0.003$ , significantly predicted counterfactual performance (figure 4b), with a marginal effect of age,  $F_{1, 155} = 2.77$ ,  $p = 0.098$ . There was no interaction between cohort and pretend performance,  $F_{2, 155} = 0.12$ ,  $p = 0.88$ . None of the other predictors or interactions approached significance,  $F \leq 1.76$ ,  $p \geq 0.18$ .

In order to examine whether there is a difference in the strength of the relationship between pretend effects score and counterfactual effects score between any of the three cohorts, we compared the slope of the relationship between pretend effects and counterfactual effects scores for all pairs of cohorts. Paired contrasts between cohorts revealed no differences between any pairs of cohorts in the slope of the relationship between pretend effects and counterfactual effects performance, mixed-SES versus low-SES U.S.,  $t_{155} = 0.38$ ,  $p = 0.70$ , mixed-SES versus Peru,  $t_{155} = 0.09$ ,  $p = 0.93$ , low-SES versus Peru,  $t_{155} = 0.47$ ,  $p = 0.64$  (figure 4b). Overall, the relationship between the two measures was the same across cohorts.

Finally, we conducted an exploratory analysis to examine the relationship between pretense and counterfactual performance within each cohort. The relationship was significant in the middle class U.S.,  $t_{173} = 3.29$ ,  $p = 0.001$ , and Peruvian  $t_{173} = 2.92$ ,  $p = 0.004$  cohorts, with a marginal trend in the same direction for the low-SES sample  $t_{173} = 1.71$ ,  $p = 0.089$ . This relationship also did not change when only children who identified the zando correctly on the first try during the initial causal scenario were included, middle class U.S.,  $t_{152} = 2.77$ ,  $p = 0.006$ , and Peruvian  $t_{152} = 2.86$ ,  $p = 0.005$  cohorts, low-SES sample  $t_{152} = 1.89$ ,  $p = 0.060$ .

Like the mixed-SES U.S. sample from Buchsbaum *et al.* [13], Peruvian and low-SES U.S. children's ability to consider counterfactual questions about the causal system was predicted by their ability to reason about the causal system during pretense. Pretend effects score significantly predicted counterfactual effects score, and age, general cognitive development as measured by conservation of number, and differences in children's inhibitory control did not account for the relationship. It is worth noting that, when the cohorts were examined individually, the correlation was significant

505 within the mixed-SES and Peruvian cohorts, but marginal in  
 506 the low-SES U.S. cohort, and we return to this point in the  
 507 General Discussion. However, we found no interaction  
 508 between cohort and pretense score in predicting counterfac-  
 509 tual performance, and no differences between any pair of  
 510 cohorts in the slope of the relationship between pretense  
 511 and counterfactual performance, suggesting that the pretend  
 512 effects score is equally predictive of counterfactual effects  
 513 score across all three cohorts (figure 3). Across cohorts, chil-  
 514 dren who responded consistently to the pretense questions  
 515 also scored higher on consistency for the counterfactual ques-  
 516 tions. Taken together, this indicates that the relationship  
 517 between pretense and counterfactual reasoning is similar  
 518 across cultures, even though children's average ability to  
 519 answer these questions differed across cultures.

## 522 7. Secondary tasks

### 524 (a) Day night stroop task

525 Children's performance on the day/night inhibition task was  
 526 significantly predicted by age and did not differ across  
 527 cohorts, see electronic supplementary material for additional  
 528 details.

### 531 (b) Conservation task

532 Children's performance on the conservation task was signifi-  
 533 cantly predicted by age, and also by cohort, with the low-SES  
 534 U.S. cohort outperforming the other two cohorts. See  
 535 supplementary material for additional details.

### 537 (c) Open-ended pretense task

538 At the end, a majority of children in the Peruvian (77%) and  
 539 low-SES U.S. (81%) cohorts suggested at least one additional  
 540 scenario to pretend with the experimenter (this measure was  
 541 not included in Buchsbaum *et al.*'s [13] Experiment 2, but was  
 542 included in Experiment 1 with the same population (71%)).  
 543 There was no difference between the three cohorts, Fisher's  
 544 Exact Test,  $p = 0.38$ . This supports the notion that on average  
 545 children were engaged in the pretend premise and activity.  
 546 See electronic supplementary material for additional details.

## 549 8. General discussion

551 This study is the first to examine the relationship between  
 552 causal counterfactual reasoning and pretense outside of a  
 553 U.S. middle class sample. Across samples, we observed a  
 554 relationship between the development of causal reasoning  
 555 during pretense and causal counterfactual reasoning, sup-  
 556 porting our primary hypothesis that these two skill sets  
 557 may be related. Additionally, we saw that, across the age  
 558 range tested, Peruvian and low-SES children scored lower  
 559 on measures of counterfactual reasoning and pretense than  
 560 did middle class U.S. children, supporting the possibility  
 561 that sociocultural variation in pretense activities may influ-  
 562 ence the development of abstract causal reasoning more  
 563 generally. Third, we found higher levels of competence in  
 564 the pretense-based causal reasoning task than we did in the  
 565 reality-based counterfactual reasoning task, particularly in  
 566 the Peruvian and low-SES U.S. cohorts, supporting the possi-  
 567 bility that earlier emerging pretense abilities may facilitate

later developing counterfactual reasoning abilities. Each of  
 these findings will be further discussed below.

Across samples, children who provided answers during  
 pretense that were consistent with the real-world causal  
 relationship were also more likely to provide accurate  
 responses to the counterfactual questions. The relationship  
 between performance on the pretend effects tasks and coun-  
 terfactual tasks held even when age, day/night score and  
 conservation score are accounted for. This replicates Buchs-  
 baum *et al.*'s [13] finding with middle class U.S. children  
 and further supports our primary hypothesis that pretense  
 and counterfactual reasoning draw upon the same emerging  
 cognitive capacities.

These findings are particularly remarkable given our  
 second finding, the observed differences across samples: the  
 scores for both the counterfactual and pretense causal reason-  
 ing questions did differ across samples, with mixed-income  
 U.S. children scoring more causally consistently than Peru-  
 vian and low-SES U.S. children on both counterfactual and  
 pretense questions. Moreover, and strikingly, this did not  
 hold for the secondary executive function and conservation  
 measures. The divergence between the performance on the  
 secondary tasks and the pretense and counterfactual tasks  
 suggests that the cohort differences were not simply the  
 result of differences in general cognitive development, or in  
 engagement or motivation to participate.

We also observed that children, particularly those in the  
 low-SES U.S. and Peruvian cohorts, provided more causally  
 consistent responses to the pretense questions than to the  
 counterfactual questions, suggesting that they may have  
 found it easier to reason about causal scenarios when they  
 were presented as pretense than when they were presented  
 as hypotheticals. In fact, Peruvian and low-SES U.S. children  
 provided causally consistent responses at above chance levels  
 only on the pretense questions, but scored at chance on the  
 counterfactual questions.

Other researchers have found that pretense may enhance  
 young children's ability to reason hypothetically [24–27]. For  
 example, Dias & Harris [24] suggested that the fantasy con-  
 text may enhance children's ability to 'quarantine' a given  
 false premise from real-world knowledge, which then  
 allows children to make more accurate downstream infer-  
 ences. Our findings are in line with these claims. By  
 contrast, the mixed-SES U.S. children did not differ signifi-  
 cantly in their pretense and counterfactual performance,  
 perhaps because they were already quite successful at both  
 tasks, and close to ceiling at the older end of our age range.  
 As this group may have experienced more opportunity to  
 engage in pretend play early in development, it is possible  
 that these differences in performance are a result of the cul-  
 tural and socioeconomic environment's impact on pretend  
 play, which we will return to below.

These findings, along with the general developmental  
 timeline in which pretend play and counterfactual reasoning  
 emerge, could imply that the ability to reason hypothetically  
 about alternatives emerges first during pretense, and reality-  
 based hypothetical reasoning emerges later. This raises a  
 compelling question about causation versus correlation:  
 might practicing causal reasoning during pretense actually  
 enhance counterfactual reasoning skills? While our findings  
 are supportive of this possibility, there are reasons to caution  
 against such a strong claim without follow-up research; most  
 importantly, our study design was correlational and not

intended to answer questions about causation. In addition, these results apply to the broader definition of counterfactuals we described in the introduction. Whether pretense correlates with the more specifically defined counterfactuals in other accounts, such as reasoning about nearest possible alternatives to past events, is an interesting open question.

As noted above, the magnitude of the correlation between pretend effects score and counterfactual effects score did not significantly differ across our populations. We found no interaction ( $p=0.88$ ) and no differences between any of the groups in the slope of the relationship between pretense performance and counterfactual performance ( $p \geq 0.64$ ). However, when the cohorts were examined individually, the relationship between these variables in the low-SES U.S. cohort was marginal. Given the otherwise statistically and qualitatively similar relationship across cohorts, this may be due to the reduction in power when examining the cohorts individually, or due to relatively increased variability within the low-SES U.S. sample. However, we also cannot rule out the possibility of a subtler systematic difference in the strength of the relationship between these variables across cohorts, which our study was not powered to detect. The question of whether not only the existence but the strength of this correlation is consistent across populations is an interesting one and could be pursued in future work with larger samples per cohort.

Further, in order to avoid children responding with respect to a familiar causal schema or script, and to directly compare pretense and counterfactual reasoning skills while controlling for content, our study introduced children to a novel causal system, and then asked them pretense questions about this same system. The similarity between the pretense and counterfactual measures allowed us to directly compare children's performance when asked similar questions about this novel causal system framed hypothetically versus when presented in a pretense scenario. However, this result does not allow us to speak to the generality of this relationship between pretense and counterfactual reasoning, and whether or not pretense writ large may facilitate counterfactual reasoning (a possibility we discuss further below). A longitudinal study, collecting a naturalistic measure of children's pretense activities, and measuring that against counterfactual reasoning performance would be a particularly effective way to address this causal question in future work.

In regard to the difference across samples, the cultural and socioeconomic environment could shape the development of pretense and counterfactual reasoning in a number of ways. For example, children who spend more time pretending may show advanced causal reasoning abilities while pretending and reasoning counterfactually, simply because they have the advantage of increased practice. Furthermore, children from families who value pretend more may spend more time pretending because they are encouraged by adults to do so. However, culture or SES may shape the type of pretense children engage in as well. Engaging in pretense with parents or older children, for example, may scaffold the complexity of children's pretense, and encourage them to think about more complex and varied types of causal relationships. Older children and adults may also correct or shape the inferences that children make during pretense and guide them towards reasoning about relationships that are causally consistent with reality. Of course, it is possible that some other variable accounts for

the differences witnessed across samples; this is a question for future research.

The observed cognitive differences across samples, notably, were specific to the counterfactual and pretense causal reasoning tasks. In fact, when contrasting across samples on the available measures, we saw more patterns of similarities than differences. First, we saw that children across samples were equally able to learn the basic real-world cause and effect relationship via observation: that a zando activated the machine, while a non-zando did not. Further, Wente *et al.* [46] conducted a different 'blicket machine' physical causal reasoning experiment with comparable samples of Peruvian, low-income U.S. children and mixed-income U.S. children. In this experiment, all children were equally able to learn much more complex physical causal relationships via observation. This may imply that all children are equally skilled at learning causal structure through observation, and socioeconomic differences are specific to the development of counterfactual causal reasoning skills. This could be further explored through research.

Relatedly, we saw comparable performance across samples on the day/night stroop task and conservation of numbers task. In another study, Zhao *et al.* [47] also found comparability across a similar sample of Peruvian and middle class U.S. children when leveraging an identical conservation of numbers task and executive functioning task. This again, suggests comparability across Peruvian and mixed-SES U.S. samples in these measured facets of broader cognitive development.

The children from Head Start, on the other hand, scored somewhat more accurately than the children from the middle class U.S. and Peruvian backgrounds on the conservation of number task, although comparably on the executive functioning task. One reason for the conservation findings may be the slight difference in ages, with the low-income Head Start children averaging older than the middle class U.S. children; however, replication would be useful before making further claims as these results were unexpected.

Our study design leveraged very similar verbal questions and procedures across pretense and counterfactual reasoning measurements. While this allows us to provide regularities in measurement across skill sets, it may also increase the concern that individual differences in verbal ability could account for at least some of the findings, such as the correlation found across measures. While the present study did control for several aspects of development, including age, conservation of number understanding and executive functioning, it did not explicitly control for verbal abilities (which were also not explicitly measured for the [13] sample).

Although we cannot rule out a possible mediating effect of verbal ability on the relationship between pretense and counterfactual reasoning task performance, previous work has also found a relationship between performance on the day/night task and verbal ability, in children in general and specifically within a HeadStart population [48]. However, in our study, performance on this task did not differ across samples and did not correlate with pretense or counterfactual task performance, suggesting that a broad effect of verbal ability on performance is less likely to explain the differences in performance across the counterfactual and pretense tasks. Nonetheless, it is possible that other common factors, such as verbal or non-verbal IQ, might underpin

the correlation. Measures of verbal ability and IQ would be valuable additional factors to consider in future work.

## 9. Conclusion

Findings provide support for the hypothesis that causal reasoning during pretend play is related to, and perhaps easier for children than causal counterfactual reasoning. Additionally, findings suggest that the environmental context impacts the development of both types of reasoning. Even so, the relationship between counterfactual reasoning and causal reasoning while pretending was very similar across samples. This suggests that, even if developmental timelines vary by environment, this relationship nonetheless remains stable. Moreover, there has been some debate about benefits of pretend play for other types of cognitive capacities (see e.g. [9]). These findings at least suggest that interventions that encourage pretend play might facilitate counterfactual reasoning, in particular, though this possibility has yet to be tested.

**Ethics.** The research protocol was approved by the Institutional Review Boards at the University of California Berkeley, protocol numbers 2010-01-631, 2015-03-7243; as well as at Pontificia Universidad Católica del Perú, protocol number N°013-2017/CEI-PUCP. Parents provided written consent prior to their children participating, and children provided verbal assent prior to participating.

**Data accessibility.** All anonymized and coded data files used for the statistical analyses in this project, as well as the R analysis scripts are available here: <https://osf.io/2udb5/>.

The data are provided in the electronic supplementary material [49].

**Authors' contributions.** A.W.: conceptualization, formal analysis, investigation, methodology, writing—original draft and writing—review and editing; A.G.: conceptualization, funding acquisition, resources, supervision and writing—review and editing; M.F.F.: methodology, resources and validation; T.G.: data curation, investigation and project administration; D.B.: conceptualization, formal analysis, methodology, supervision, visualization, writing—original draft and writing—review and editing.

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## EndNotes

<sup>1</sup>We consider pretend play to be play in which the child acts on or in relation to a transformed or alternative reality, intentionally acting 'as if' a non-real state of the world were true [1,2].

<sup>2</sup>Counterfactual reasoning is a wide-ranging interdisciplinary research area. While some researchers use the term 'counterfactual reasoning' to refer to any reasoning about alternative representations that conflict with actual representations of reality, others use the term to refer specifically to reasoning about nearest possible alternatives to past events (see e.g. [7,8]; for some discussion). In this work, we use the broader definition of counterfactual reasoning, meaning any reasoning about alternatives that conflict with reality.

<sup>3</sup>These questions required children to reason about the future effects of present counterfactual identities, in contrast with some previous work which has asked children to reason about the effect of counterfactual past actions on present events.

<sup>4</sup>The counterfactual effects and pretend effects questions directly parallel each other—one asks about the counterfactual causal effect of the block given a counterfactual identity, and one asks about the pretend causal effect of the block given a pretend identity. Therefore, as in [13], we focus our analyses on pretend effects performance. However, see electronic supplementary material for an analysis of the relationship between the pretend intervention performance and counterfactual effects performance.

<sup>5</sup>One child in Peru, four in the low-SES U.S. sample, and two in the Buchsbaum *et al.* [13] sample did not complete one or both of the secondary tasks, and so their data are not included in analyses using these measures.

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